



# Probability, Third Edition

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## Solutions to practice questions – Chapter 9

### Solution 9.1

Let  $X$  be the number of demands made each day, and let  $Y$  be the number of demands handled.

Then we have:

$$Y = \min\{X, 3\}$$

and:

$$\Pr(Y = 0) = \Pr(X = 0) = e^{-2} = 0.13534$$

$$\Pr(Y = 1) = \Pr(X = 1) = 2e^{-2} = 0.27067$$

$$\Pr(Y = 2) = \Pr(X = 2) = \frac{2^2 e^{-2}}{2!} = 0.27067$$

$$\Pr(Y = 3) = \Pr(X \geq 3) = 1 - 0.13534 - 0.27067 - 0.27067 = 0.32332$$

Hence, the expected number of demands handled is:

$$\begin{aligned} E[Y] &= 0 \times 0.13534 + 1 \times 0.27067 + 2 \times 0.27067 + 3 \times 0.32332 \\ &= 1.782 \end{aligned}$$

### Solution 9.2

We can calculate  $E[Y]$  and  $E[Y^2]$  in terms of  $E[X]$  and  $E[X^2]$ :

$$E[Y] = \sum_{y=0}^{\infty} y \Pr(Y = y) = \sum_{y=1}^{\infty} y \Pr(Y = y) = (1 - \alpha) \sum_{y=1}^{\infty} y \Pr(X = y) = (1 - \alpha) E[X] = (1 - \alpha) \mu$$

$$E[Y^2] = \sum_{y=0}^{\infty} y^2 \Pr(Y = y) = \sum_{y=1}^{\infty} y^2 \Pr(Y = y) = (1 - \alpha) \sum_{y=1}^{\infty} y^2 \Pr(X = y) = (1 - \alpha) E[X^2] = (1 - \alpha)(\mu^2 + \mu)$$

Hence, the variance of  $Y$  is:

$$\begin{aligned} \text{var}(Y) &= E[Y^2] - (E[Y])^2 \\ &= (1 - \alpha)(\mu^2 + \mu) - ((1 - \alpha)\mu)^2 \\ &= (1 - \alpha)\mu(1 + \alpha\mu) \end{aligned}$$

**Solution 9.3**

Note that since the pdf is defined for  $x > 0$ , the transformation  $Y = X^2$  is 1-1, with:

$$X = +\sqrt{Y}$$

So, using the method of transformations:

$$f_Y(y) = f_X(\sqrt{y}) \left| \frac{d(\sqrt{y})}{dy} \right| = e^{-\sqrt{y}} \times \frac{1}{2\sqrt{y}} \quad \text{for } y > 0$$

**Solution 9.4**

The pdf of  $X$  is:

$$f_X(x) = \frac{x^{\alpha-1} e^{-x/\theta}}{\theta^\alpha \Gamma(\alpha)} \quad \text{for } x > 0$$

Now, if  $Y = g(X) = cX$ , we have a 1-1 differentiable transformation.

The inverse transformation is:

$$X = Y/c$$

So, using the method of transformations:

$$\begin{aligned} f_Y(y) &= f_X\left(\frac{y}{c}\right) \left| \frac{d(y/c)}{dy} \right| = \frac{(y/c)^{\alpha-1} e^{-(y/c)/\theta}}{\theta^\alpha \Gamma(\alpha)} \times \frac{1}{c} \\ &= \frac{y^{\alpha-1} e^{-y/(c\theta)}}{(c\theta)^\alpha \Gamma(\alpha)} \quad \text{for } y > 0 \end{aligned}$$

Hence  $Y$  follows a gamma distribution with parameters  $\alpha$  and  $c\theta$ .

**Solution 9.5**

Let's use the method of distribution functions.

We have:

$$F_Y(y) = \Pr(Y \leq y) = \Pr\left(\frac{1}{X} - 1 \leq y\right) = \Pr\left(X \geq \frac{1}{1+y}\right) = 1 - \Pr\left(X < \frac{1}{1+y}\right) = 1 - \frac{1}{1+y}$$

Hence, the pdf of  $Y$  is:

$$f_Y(y) = F'_Y(y) = \frac{1}{(1+y)^2} \quad \text{for } y > 0$$

**Solution 9.6**

The density of  $T$  is:

$$f(t) = \frac{1}{3}e^{-t/3} \quad \text{for } t > 0$$

The discovery time is a piecewise function of  $T$ :

$$X = g(T) = \max(T, 2) = \begin{cases} 2 & T \leq 2 \\ T & T > 2 \end{cases}$$

Hence:

$$\begin{aligned} E[X] &= \int_0^{\infty} g(t)f(t)dt = 2\int_0^2 \frac{e^{-t/3}}{3} dt + \int_2^{\infty} \frac{te^{-t/3}}{3} dt \\ &= 2 \times F(2) + \left(-te^{-t/3} - 3e^{-t/3}\right) \Big|_2^{\infty} \quad (\text{integrating by parts}) \\ &= 2(1 - e^{-2/3}) + (-0 - 0 + 2e^{-2/3} + 3e^{-2/3}) \\ &= 2 + 3e^{-2/3} \end{aligned}$$

**Solution 9.7**

We are given that  $T$  is distributed uniformly on the interval  $[8, 12]$ , so the probability density function is:

$$f_T(t) = \frac{1}{4} \quad \text{for } 8 \leq t \leq 12$$

The random rate of service in customers per minute is:

$$R = 10/T$$

Since this is a 1-1, differentiable transformation of  $T$  with inverse function  $T = 10/R$ , we can calculate the probability density function  $f_R(r)$  using the method of transformations:

$$\begin{aligned} f_R(r) &= f_T(g^{-1}(r)) \left| \frac{dg^{-1}(r)}{dr} \right| \\ &= f_T\left(\frac{10}{r}\right) \times \left| \frac{d(10/r)}{dr} \right| \\ &= \frac{1}{4} \times \left| \frac{-10}{r^2} \right| \\ &= \frac{5}{2r^2} \end{aligned}$$

**Solution 9.8**

Let  $Y$  be the length of time lived in the next 10 years.

Then  $Y$  is defined as follows:

$$Y = \min(X, 10) = \begin{cases} X & 0 \leq X \leq 10 \\ 10 & X > 10 \end{cases}$$

The expected value of  $Y$  is therefore:

$$\begin{aligned} E[Y] &= \int g(x) f(x) dx \\ &= \int_0^{10} x \frac{1}{75} dx + \int_{10}^{75} 10 \frac{1}{75} dx \\ &= \frac{x^2}{150} \Big|_0^{10} + \frac{10x}{75} \Big|_{10}^{75} \\ &= 9.333 \end{aligned}$$

**Note:** This question uses the theory related to policy limits, since we apply a maximum of 10 to the future lifetime of the 25-year-old.

**Solution 9.9**

The benefit amount paid by the insurer is:

$$Y = \max\{0, (X - 250)\} = \begin{cases} 0 & 0 \leq X \leq 250 \\ X - 250 & X > 250 \end{cases}$$

So the expected benefit amount paid by the insurer is:

$$\begin{aligned} E[Y] &= \int_0^{250} 0 \times f_X(x) dx + \int_{250}^{\infty} (x - 250) f_X(x) dx \\ &= 0 + \int_{250}^{\infty} (x - 250) \frac{1}{1,000} e^{-x/1,000} dx \end{aligned}$$

We could use integration by parts to finish the calculation. However, it is quite a bit quicker to employ the substitution  $z = x - 250$ :

$$\begin{aligned} E[Y] &= \int_0^{\infty} z \frac{1}{1,000} e^{-(z+250)/1,000} dz \\ &= e^{-250/1,000} \int_0^{\infty} z \frac{1}{1,000} e^{-z/1,000} dz \end{aligned}$$

Finally, note that the last integral is the expected value of an exponential distribution with mean 1,000. So without performing any further integration, we can see that the final result is:

$$E[Y] = e^{-0.25} \times 1,000 = 778.80$$

**Solution 9.10**

Let  $X$  be the manufacturer's annual losses, and let  $Y$  be the part of the annual losses not paid by the insurance company. Then:

$$Y = \min(X, 2) = \begin{cases} X & 0.6 < X \leq 2 \\ 2 & X > 2 \end{cases}$$

We can calculate the expected value of  $Y$  as:

$$\begin{aligned} E[Y] &= \int_{0.6}^2 x f(x) dx + \int_2^{\infty} 2 f(x) dx \\ &= 2.5(0.6)^{2.5} \int_{0.6}^2 x^{-2.5} dx + 2 \times 2.5(0.6)^{2.5} \int_2^{\infty} x^{-3.5} dx \\ &= -\frac{2.5(0.6)^{2.5}}{1.5x^{1.5}} \Big|_{0.6}^2 - \frac{2(0.6)^{2.5}}{x^{2.5}} \Big|_2^{\infty} \\ &= \frac{2.5}{1.5} \left( 0.6 - \frac{(0.6)^{2.5}}{(2)^{1.5}} \right) + 2 \left( \frac{0.6}{2} \right)^{2.5} \\ &= 0.9343 \end{aligned}$$

**Solution 9.11**

Since  $Y = 1.05X$  is a 1-1 differentiable transformation, we can use the method of transformations.

First, we identify the inverse function of the transformation. This is simply:

$$x = y/1.05$$

Hence, the density function of  $Y$ , the loss in 2003, is given by:

$$\begin{aligned} f_Y(y) &= f_X(g^{-1}(y)) \left| \frac{dg^{-1}(y)}{dy} \right| = f_X(y/1.05) \left| \frac{d(y/1.05)}{dy} \right| \\ &= \frac{2 \times 1,000^2}{((y/1.05) + 1,000)^3} \times \frac{1}{1.05} \\ &= \frac{2 \times 1,050^2}{(y + 1,050)^3} \quad \text{for } y > 0 \end{aligned}$$

**Solution 9.12**

Let  $Y = X_1 + X_2$  where  $X_1$  and  $X_2$  are the lifetimes of the primary and backup systems.

The joint density of  $X_1$  and  $X_2$  is:

$$f(x_1, x_2) = f(x_1)f(x_2) = 0.05e^{-0.05x_1} \times 0.10e^{-0.10x_2} \quad \text{where } x_1, x_2 > 0$$

Step 1: The region in the  $X_1 \times X_2$  plane defined by the inequality  $X_1 + X_2 \leq y$  is a triangular region in the first quadrant below the line  $X_1 + X_2 = y$ .

Step 2: The cdf of  $Y$  is calculated as:

$$\begin{aligned} F_Y(y) &= \Pr(X_1 + X_2 \leq y) = \int_{x_2=0}^y \int_{x_1=0}^{y-x_2} 0.05e^{-0.05x_1} 0.10e^{-0.10x_2} dx_1 dx_2 \\ &= \int_{x_2=0}^y 0.10e^{-0.10x_2} \left(1 - e^{-0.05(y-x_2)}\right) dx_2 \\ &= \int_0^y 0.10e^{-0.10x_2} dx_2 - 2e^{-0.05y} \int_0^y 0.05e^{-0.05x_2} dx_2 \\ &= \left(1 - e^{-0.10y}\right) - 2e^{-0.05y} \left(1 - e^{-0.05y}\right) \\ &= 1 + e^{-0.10y} - 2e^{-0.05y} \end{aligned}$$

Step 3: Differentiate this to obtain the density function of  $Y$ :

$$f_Y(y) = F'_Y(y) = 0.1(e^{-0.05y} - e^{-0.10y})$$

**Solution 9.13**

⇒ From 2010, this question is no longer covered by the Exam P syllabus.

**Solution 9.14**

For an exponential distribution with mean 10, we have:

$$f(x) = 0.1e^{-0.1x}$$

$$F(x) = 1 - e^{-0.1x}$$

Hence, the pdf of the 4th largest claim (which is also the 2nd smallest claim) is:

$$\begin{aligned} f_{Y_2}(y) &= \frac{5!}{1!3!} (F_X(y))^1 (1 - F_X(y))^3 f_X(y) \\ &= 20 \times (1 - e^{-0.1y}) \times (e^{-0.1y})^3 \times 0.1e^{-0.1y} \\ &= 2(e^{-0.4y} - e^{-0.5y}) \end{aligned}$$

Hence the expected value is:

$$\begin{aligned} \int_0^{\infty} y f_{Y_2}(y) dy &= \int_0^{\infty} 2y(e^{-0.4y} - e^{-0.5y}) dy \\ &= 5 \int_0^{\infty} 0.4ye^{-0.4y} dy - 4 \int_0^{\infty} 0.5ye^{-0.5y} dy \end{aligned}$$

Notice that these two integrals are the expected values of exponential distributions with means 2.5 and 2, hence:

$$E[Y_2] = 5 \times 2.5 - 4 \times 2 = 4.5$$

**Solution 9.15**

The first step here is to calculate the density function of:

$$Y = \max\{X_1, X_2, X_3\}$$

We can use the method of distribution functions to calculate the distribution function of  $Y$  as follows:

$$\begin{aligned} F_Y(y) &= \Pr(Y \leq y) = \Pr(\text{all } X_i \leq y) = (F_X(y))^3 \\ &= \left( \int_1^y \frac{3}{x^4} dx \right)^3 = \left( 1 - \frac{1}{y^3} \right)^3 \quad y > 1 \end{aligned}$$

We differentiate the cdf to find the pdf:

$$f_Y(y) = F'_Y(y) = 3 \left( 1 - \frac{1}{y^3} \right)^2 \frac{3}{y^4} = 9 \left( \frac{1}{y^4} - \frac{2}{y^7} + \frac{1}{y^{10}} \right) \quad y > 1$$

and then compute  $E[Y]$  as:

$$\begin{aligned} E[Y] &= \int_1^{\infty} y f_Y(y) dy = 9 \int_1^{\infty} \frac{1}{y^3} - \frac{2}{y^6} + \frac{1}{y^9} dy \\ &= 9 \left( -\frac{1}{2y^2} + \frac{2}{5y^5} - \frac{1}{8y^8} \right) \Bigg|_1^{\infty} = 9 \left( \frac{1}{2} - \frac{2}{5} + \frac{1}{8} \right) = 2.025 \text{ (thousand)} \end{aligned}$$